Assessing Mortality Risk of Climate Change

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21 March 2024



Acknowledgment

This research is sponsored by the **SCOR Foundation for Science** through the project titled "Who, When, and Where? Assessing Mortality Risk of Climate Change".



Outline of the presentation

Background and motivation

2 Data collection, processing, and visualization

Modeling framework

Empirical results



Outline

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- 3 Modeling framework
- 4 Empirical results



Climate-driven Mortality Risk

The link between climate change and human mortality

According to the WHO:

Between **2030** and **2050**, climate change is expected to cause approximately 250,000 additional deaths per vear.

Weather-related catastrophes

Extreme temperatures

Climate-sensitive infectious diseases



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How extreme temperatures affect mortality

- Globally, approximately **5 million** deaths were associated with non-optimal temperatures per year, that is **1** out of every **10** deaths in the world during 2000–2019.
- Extreme **heat** associated excess mortality is more **immediate**, while the increase in mortality following extreme **cold** is long **lasting**.
- The **elderly** are more **fragile** to extreme temperatures, and different **regions** react to extreme temperatures **differently**.







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5/46

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Our project is expected to provide insights into the following questions:

- **Who** are the excess deaths? *Find the age groups particularly sensitive to climate change.*
- **When** do excess deaths occur? *Determine if more excess deaths occur in winter or summer.*
- Where are the excess deaths? *Identify regions that are most vulnerable to climate change.*



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General trends and patterns in mortality rates



Figure 1: National-level male mortality for US: 1969–2019.



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Regional differences in mortality experience



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Geographical resolutions: continental level

Li, H., Tang, Q., 2022. Joint extremes in temperature and mortality: A bivariate POT approach. *North American Actuarial Journal*, 26(1), 43–63.



Source: Actuaries Climate Index Executive Summary (2018), Page 4, Figure 2



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Geographical resolutions: city level

Gasparrini, A., Armstrong, B., & Kenward, M. G., 2010. Distributed lag non-linear models. Statistics in Medicine, 29(21), 2224–2234.



Source: GISGeography, https://gisgeography.com/new-york-city-map



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Geographical resolutions: climate division level

Vose, Russell S., et al., 2014. Improved historical temperature and precipitation time series for US climate divisions. *Journal of Applied Meteorology and Climatology*, 53(5), 1232–1251.



Source: NOAA, https://www.ncei.noaa.gov



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Death data

The US county-level monthly death data for the time period 1968–2018 are obtained from two main sources listed as follows.

- The National Center for Health Statistics (NCHS)
- Centers for Disease Control and Prevention (CDC) WONDER

Note that:

1) As there is no publicly available data on monthly age-specific population exposure, particularly at the county level, we assume constant exposure throughout the year.

2) The CDC WONDER online database censors death counts between 0 and 9 due to privacy constraints for years 1989 onward.



A map of US counties



Source: Wikipedia, https://en.wikipedia.org/wiki/County_(United_States)



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We collect the corresponding monthly temperature data at the climate division level, from the National Oceanic and Atmospheric Administration (NOAA) for the contiguous United States.

• NOAA Monthly U.S. Climate Divisional Database

Note that:

- 1) There are 344 climate divisions, and they are larger than counties but smaller than states.
- 2) The borders of climate divisions overlap those of counties, some care will need to be taken to combine the deaths database with the weather database.



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A map of US climate divisions



Source: NOAA, https://www.ncei.noaa.gov/access/monitoring/dyk/us-climate-divisions



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Matching counties into climate divisions



Pink line: Climate division border

Grey line: County border

Blue dot: County population center



A quick glance at the data



Figure plots log mortality, ages 75–84, January 2018.



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Quantifying mortality risk due to temperatures is challenging:

- The relationship is non-linear;
- This non-linear relationship may also vary with geography; 2
- There is a need to control for time trend and age patterns that can confound the relationship. 3



The model

We propose the following model:

$$y_{i,x,t} = \mu_{i,x,t} + \varepsilon_{i,x,t} \quad \varepsilon \stackrel{i.i.d}{\sim} N(0, \sigma^2), \tag{1}$$

$$\mu_{i,x,t} = \alpha + \gamma_i^{(cd)} + \gamma_{\lfloor t/12 \rfloor}^{(yr)} + \gamma_{t \mod 12}^{(mth)} + \gamma_x^{(age)} + \gamma_{\lfloor t/12 \rfloor,x}^{(yr-age)} + f_i(\operatorname{Temp}_{i,t}),$$
(2)

where

- $y_{i,x,t}$ is (annualized) log mortality in climate division *i* at time *t* for age group *x*.
- $\gamma^{(cd)}$, γ^{yr} , γ^{mth} , γ^{age} are random effects for climate division, year, month, and age respectively.
- γ^{yr-age} is a two way random effect for year and age.
- $f_x(\text{Temp}_{i,t})$ is a climate division specific regression function measuring the effect of temperature on mortality



Data suppression

Log mortality is observed monthly but annualized

$$y_{i,x,t} = \log\left(\frac{\text{Deaths}_{i,x,t}}{\text{Population}_{i,x,t}} \times \frac{\text{Days in Year}_t}{\text{Days in Month}_t}\right)$$

When there are fewer than 10 deaths, we know that $y_{i,x,t}$ is upper bounded by

$$u_{i,x,t} = \log\left(\frac{10}{\text{Population}_{i,x,t}} \times \frac{\text{Days in Year}_t}{\text{Days in Month}_t}\right)$$



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Data augmentation

In the Bayesian framework, the model becomes

$$y_{i,x,t}^* = \boldsymbol{\mu}_{i,x,t} + \boldsymbol{\varepsilon}_{i,x,t},$$

where the latent variable y^* is defined as

$$y_{i,x,t}^* \begin{cases} = y_{i,x,t} & \text{if not suppressed} \\ \sim N(\mu_{i,x,t}, \sigma^2) I(y_{i,x,t}^* < u_{i,x,t}) & \text{if suppressed} \end{cases}$$

The latent values corresponding to suppressed observation are treated like any other parameters that can be estimated.



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Priors

• The idea is to shrink "neighboring" random effects together

- Neighboring regions;
- ▶ Neighboring years (*e.g.* 2005 and 2006);
- Neighboring months (March and April);
- ▶ Neighboring age groups (*e.g.* 55–64 and 64–75).
- This is done with an intrinsic conditional autoregressive (ICAR) prior (Besag and Kooperberg, 1995).
- Best understood with graph theory.



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ICAR prior and graphs

- A graph has nodes and edges.
 - In our case, nodes will correspond to random effects (denoted γ).
- Edges connect two nodes.
 - Edges will correspond to differences between random effects (denoted δ).
- ICAR prior assumes $\delta \sim N(0, \tau^2)$.
- Differences between neighbors shrunk to zero.
- Equivalent to assuming $\gamma \sim N(\bar{\gamma}, \tau^2)$ where $\bar{\gamma}$ is the neighborhood average.



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Climate division graph







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Age graph





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Month graph



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Two-way effects

- For two way effects construct the Cartesian product of two graphs,
- Consider age year effect,
- Nodes become Age year pairs, e.g. (2005, 55–64),
- Edge between nodes if
 - ▶ Year is the same and age are neighbours, *e.g.*, (2005,55–64) and (2005, 65–74).
 - ▶ Year are neighbours and age is the same, *e.g.*, (2005, 55–64) and (2006, 55–64).



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Two-way effects





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Regression effects for temperature modelled non-parametrically using B-Splines

$$f(\text{Temp}) = \sum_{j=1}^{p} B_j(\text{Temp})\beta_j.$$

The functions $B_j(.)$ are known as a B-Spline basis, and a known pre-computed functions, while β_j are coefficients. This makes a non-linear regression into a linear regression.



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30/46



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31/46



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33/46

Prior on B-Spline coefficients

- Use a very large number of basis functions (e.g. 20 functions).
- Penalise differences between coefficients that correspond to adjacent splines.
- Approximates a penalty on integrated squared gradient of the function.
- Used in our framework for main regression effect.
- Bayesian version of penalized splines (Eilers and Marx, 1996).



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Novelty

- We also have climate division specific temperature effects.
- The basis coefficients corresponding to the same B-spline across neighboring regions are also shrunk together (their difference is shrunk to zero).
- To the best of our knowledge this is a **novel innovation** in our work.
- Employ the same prior as for two way effects.



Recap-Priors

$$\begin{split} \delta_{(dom)} &\sim N(0, \tau^2_{(dom)}) \\ \tau_{(dom)} &\sim \text{Cauchy}(0, 1) \\ p(\alpha, \sigma) &\propto 1/\sigma^2 \end{split}$$

Here (*dom*) refers to either yearly, age, monthly, climate division, year-age effects, or B-spline basis coefficients (main and climate division specific). All γ and β are can be found deterministically from δ .



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- Inference carried out using the No U Turn Sampler (NUTS).
- The implementation is in **STAN** (Carpenter et al., 2017).
- Samples from posterior by proposing new values of the parameter vector.
- Idea is to treat posterior as a measure of energy and evolve a proposal for the parameter vector using Hamiltonian dynamics.



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Results - Spatial effect



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38/46

Results - Monthly effect



39/46

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-953

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Results - Yearly effect



40/46

Results - Age-year effect



41/46

Results - Regression



Results - Regression



Conclusions

Our insights into the following questions:

- Who are the excess deaths? *Elderly people particularly the "oldest old"*.
- When do excess deaths occur? *Extreme cold temperatures in winter months and extreme heat during summer months.*
- Where are the excess deaths? *Geographic & Socioeconomic*.



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Questions and discussions



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