

Coherent Cause-Specific Mortality Forecasting via Constrained Penalized Regression Models

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in collaboration with María Durbán (Universidad Carlos III de Madrid)

*SCOR Chair on mortality research
Workshop 1*

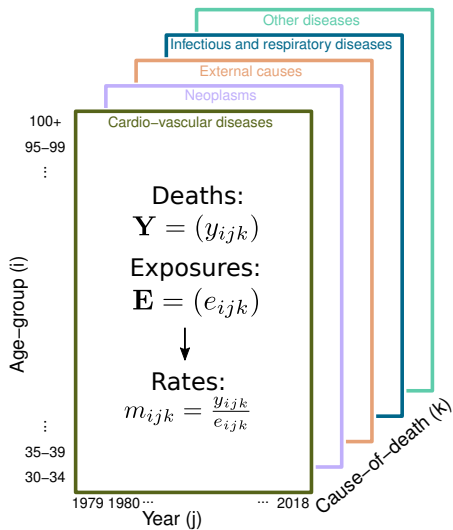
April 4-5, 2024

- Overall mortality trends are the summation of cause-specific mortality experiences

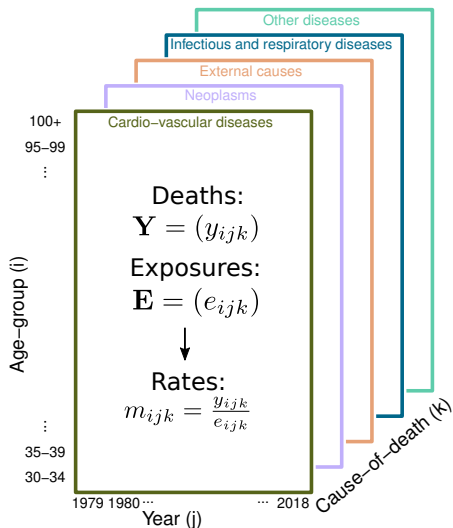
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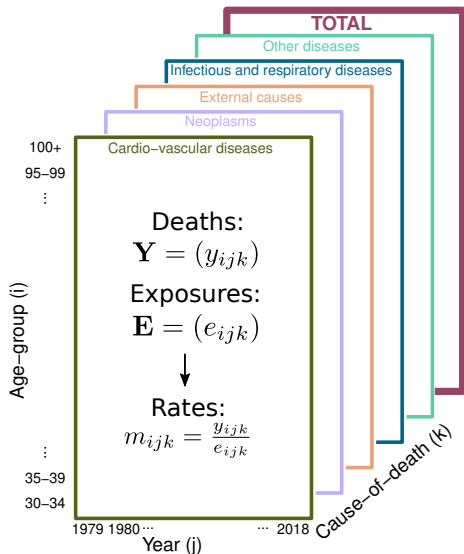
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- For long time it has been argued that all-cause mortality projections based on cause-specific mortality present serious drawbacks (Wilmoth, MPS, 1995)
- Recent developments:
 - Lee-Carter model for specific cause (Kjærgaard et al., JRSS-C, 2019)
 - Bayesian hierarchical model for cause-specific death rates in geographic subunits (Foreman et al., JRSS-C, 2017)



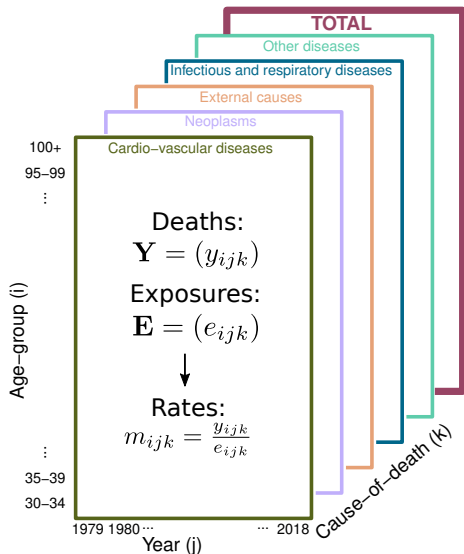
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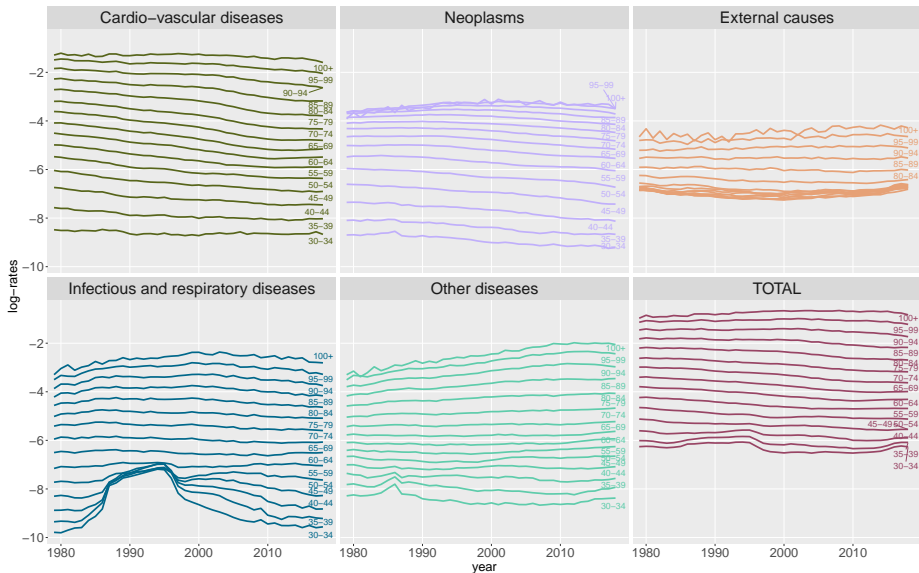


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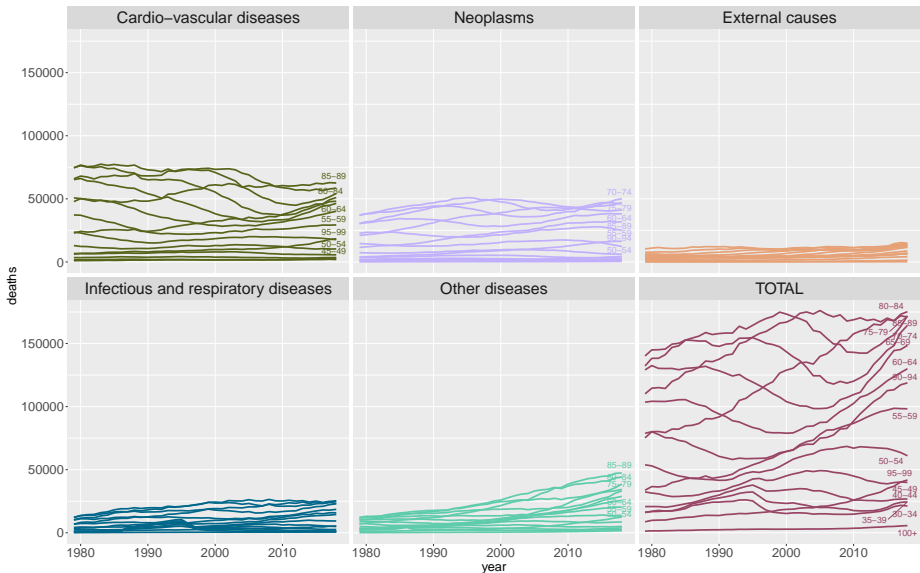


- Data are two 3-D arrays
- Classified by age (\mathbf{a}), year (\mathbf{t}) and cause of death (\mathbf{c})
- Final layer of \mathbf{Y} contains total number of deaths
- Each layer in \mathbf{E} includes the same age-year matrix

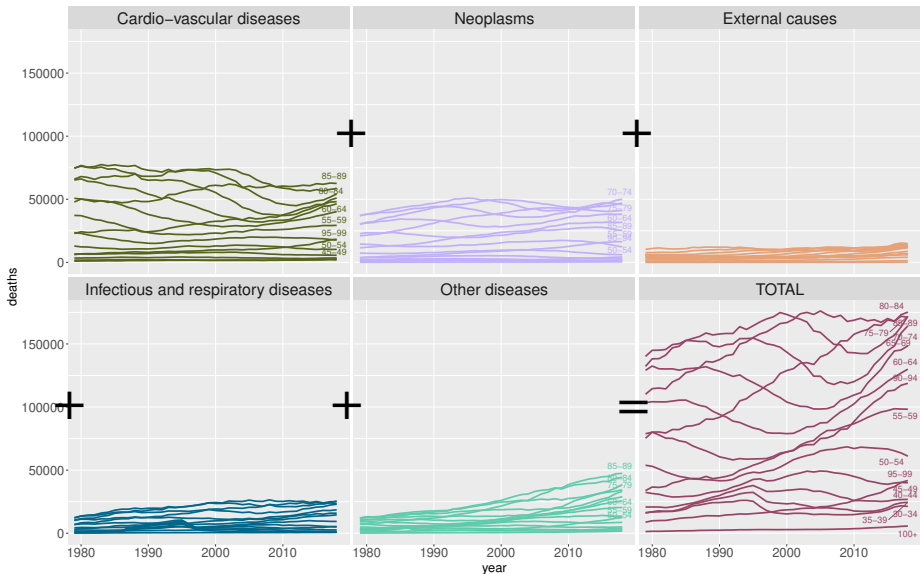
Example: USA mortality, males. Log-rates



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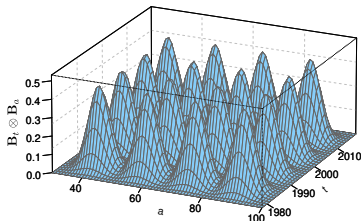
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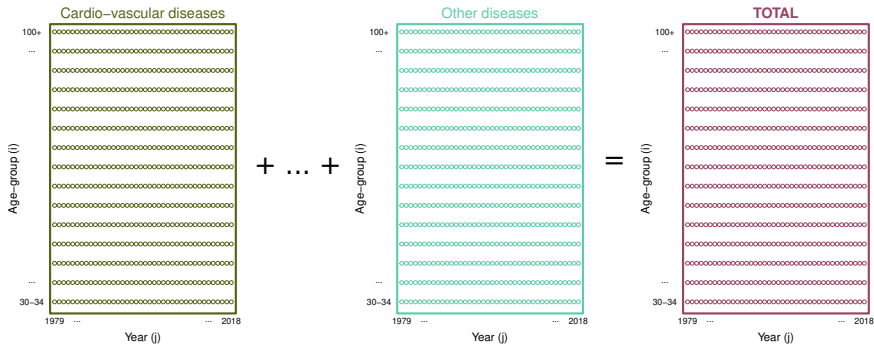
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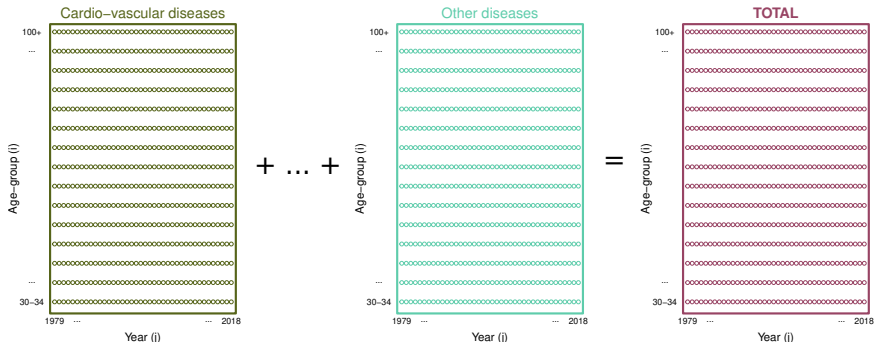
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- \mathbf{C} can be written as a Kronecker product
- \Rightarrow All inner operations can be embedded in a Generalized Linear Array Model (GLAM) framework (Currie et al., JRSS-B, 2006)

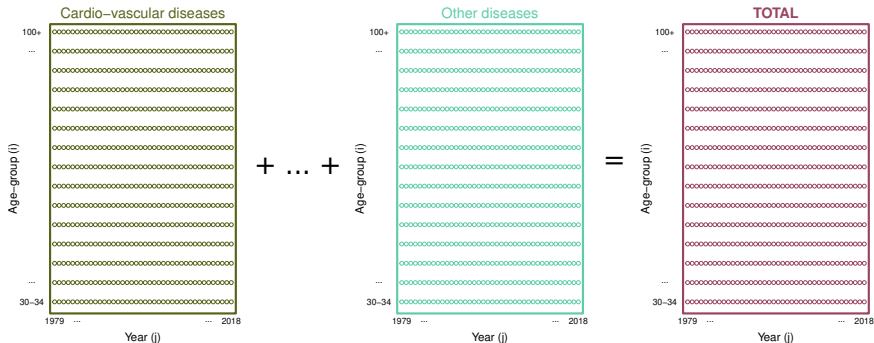
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⇒ singular system



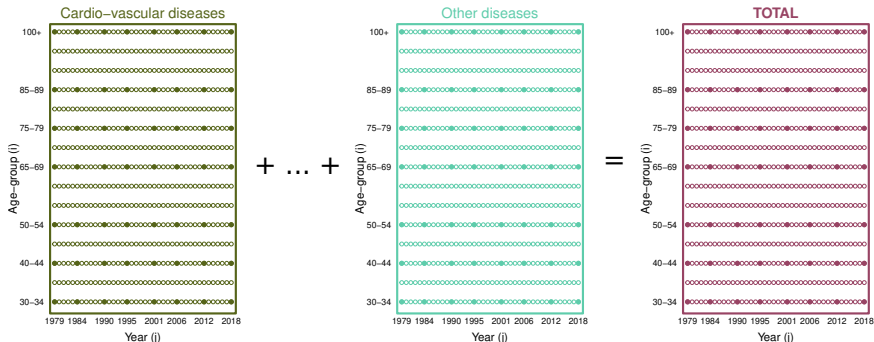
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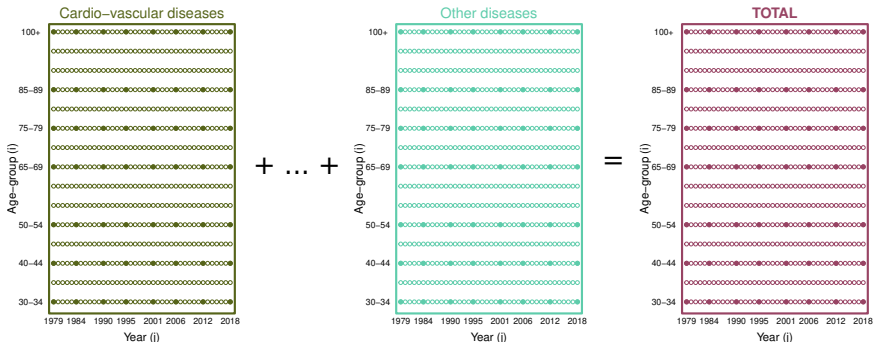
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- Enforce constraints for a number of equally-spaced data-points
- $C = C_c \otimes I_m \otimes T_n = C_c \otimes C_t \otimes C_a$



- We maximize a Poisson likelihood:

$$\ell = \mathbf{y}'\mathbf{B}\boldsymbol{\alpha} - \mathbf{e}' \exp(\mathbf{B}\boldsymbol{\alpha})$$

- We maximize a **penalized** Poisson likelihood:

$$l_P = \mathbf{y}'\mathbf{B}\boldsymbol{\alpha} - \mathbf{e}' \exp(\mathbf{B}\boldsymbol{\alpha}) - \frac{1}{2}\boldsymbol{\alpha}'\mathbf{P}\boldsymbol{\alpha}$$

- **Smoothness** over age and time for each cause is ensured by \mathbf{P} (block-diagonal structure)

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- We use smoothing parameters optimized by Bayesian Information Criterion (BIC) when estimating each cause-specific age-time matrix independently

- We maximize a **constrained penalized** Poisson likelihood:

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- **Constraints** are enforced by Lagrange multipliers $\boldsymbol{\omega}$

- To maximize

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- We derive the following scoring algorithm:

$$\begin{bmatrix} \mathbf{B}'\tilde{\mathbf{W}}\mathbf{B} + \mathbf{P} + \mathbf{B}'\text{diag}(\mathbf{C}'\tilde{\boldsymbol{\omega}})\tilde{\mathbf{V}}\mathbf{B} & \mathbf{B}'\tilde{\mathbf{V}}\mathbf{C}' \\ \mathbf{C}\tilde{\mathbf{V}}\mathbf{B} & \mathbf{0} \end{bmatrix} \begin{bmatrix} \tilde{\boldsymbol{\alpha}} \\ \tilde{\boldsymbol{\omega}} \end{bmatrix} = \begin{bmatrix} \mathbf{B}'\tilde{\mathbf{W}}\tilde{\mathbf{z}} + \mathbf{B}'\text{diag}(\mathbf{C}'\tilde{\boldsymbol{\omega}})\tilde{\mathbf{V}}\mathbf{B}\tilde{\boldsymbol{\alpha}} \\ \mathbf{C}\tilde{\mathbf{V}}\mathbf{B}\tilde{\boldsymbol{\alpha}} - \mathbf{C}\boldsymbol{\gamma} \end{bmatrix}$$

where

- $\boldsymbol{\gamma} = \exp \boldsymbol{\eta}$, $\mathbf{V} = \text{diag}(\boldsymbol{\gamma})$
- \mathbf{W} and \mathbf{z} Poisson regression weights and working response

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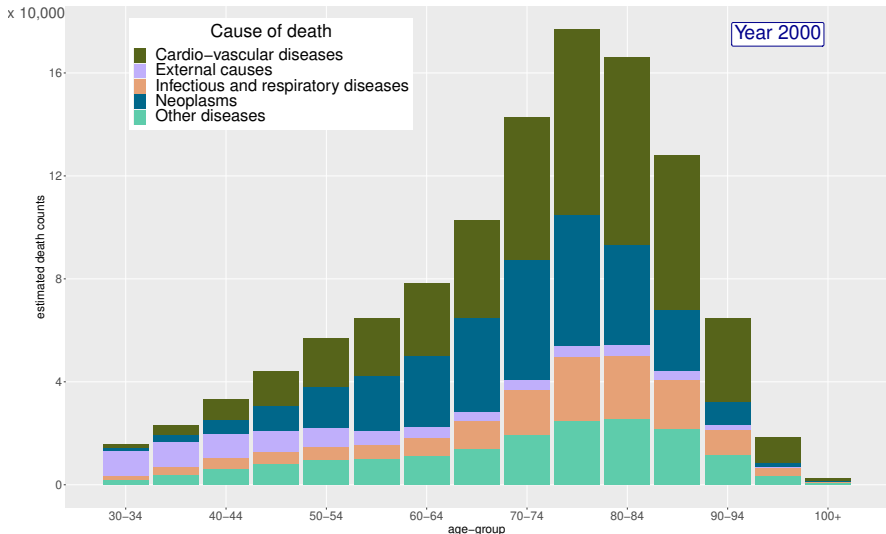
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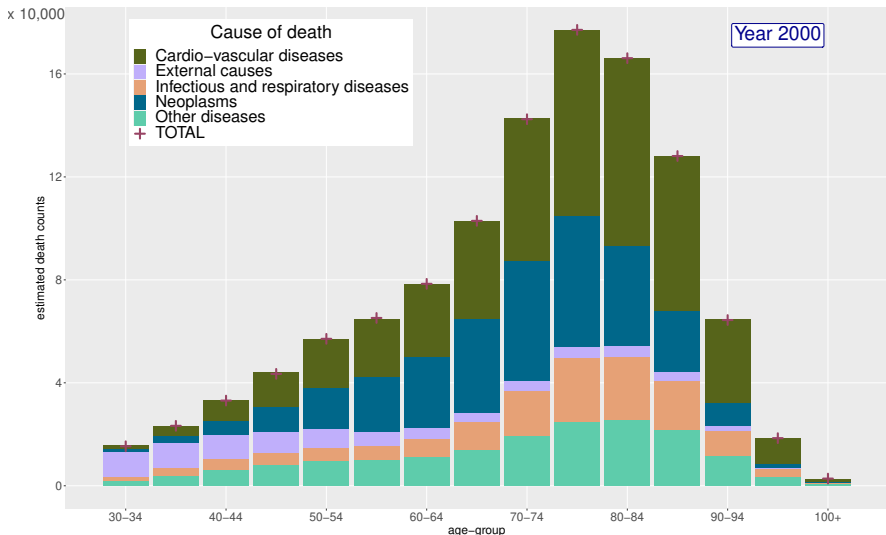
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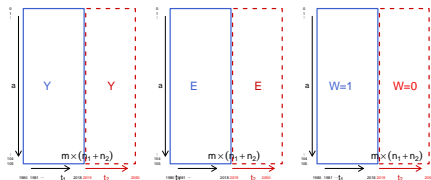
- $\boldsymbol{\gamma} = \exp \boldsymbol{\eta}$, $\mathbf{V} = \text{diag}(\boldsymbol{\gamma})$
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- With no summation constraint the model simply reduces to a series of two-dimensional GLAMs



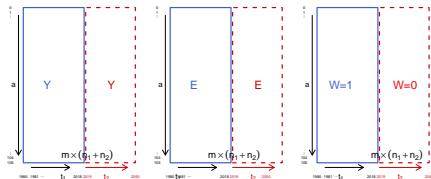


- Future is a missing value problem (Currie et al., StatMod, 2004)

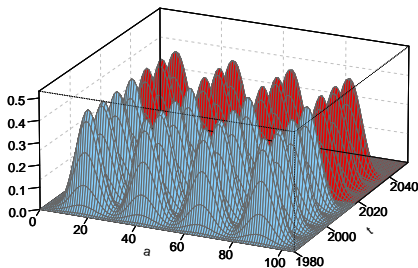
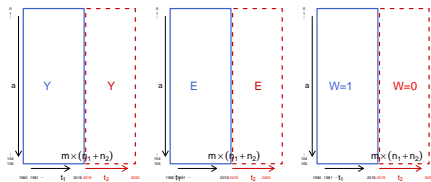
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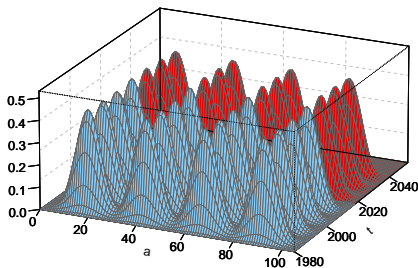
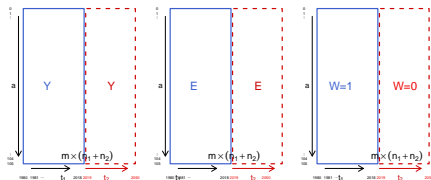
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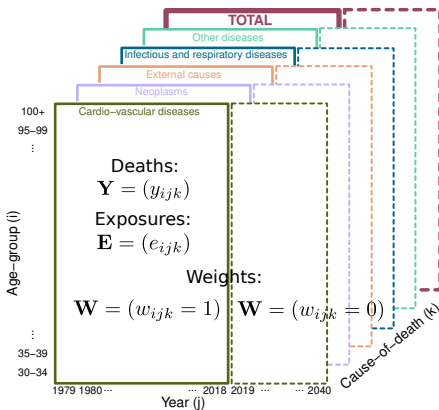
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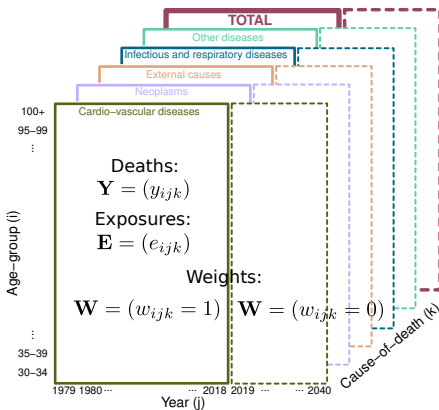
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- We add 0/1 weights in the scoring algorithm



- We augment both deaths and exposures
- We define a 0/1 weight 3D array
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- We generalize in 3D the expanded B -spline bases:

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Actual and Smooth
log-mortality

Rate-of-aging from
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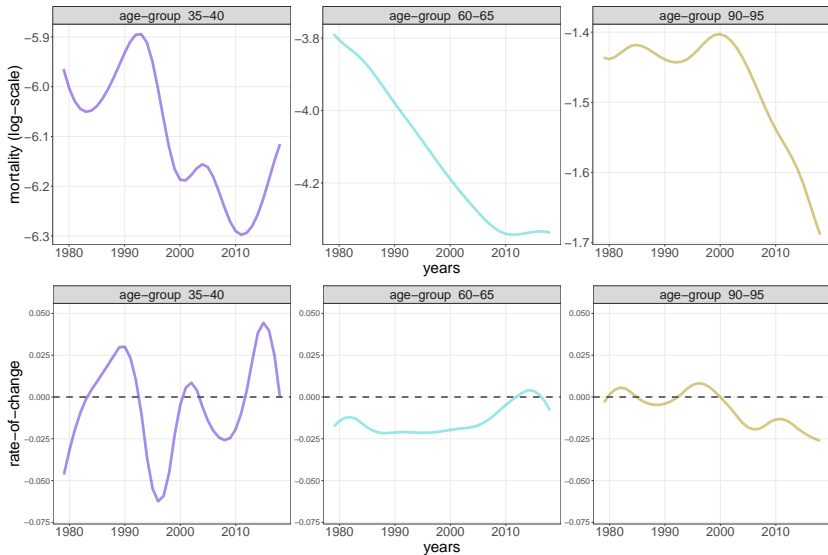
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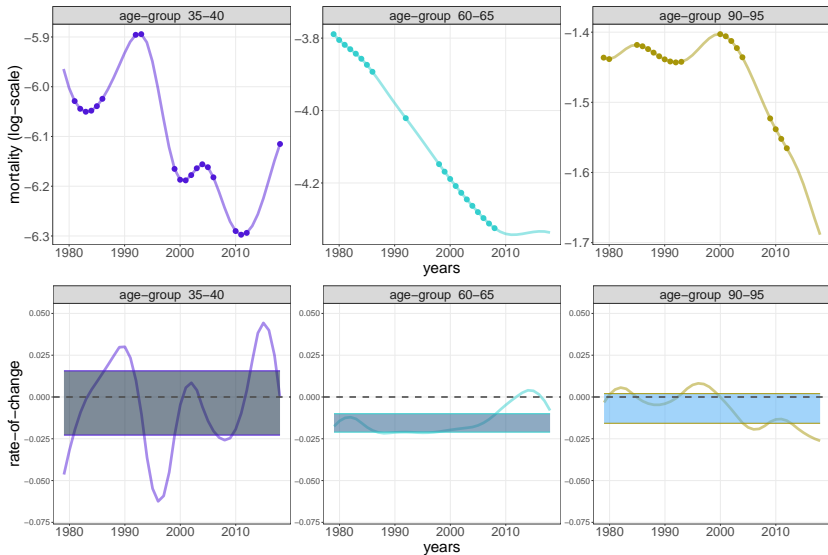
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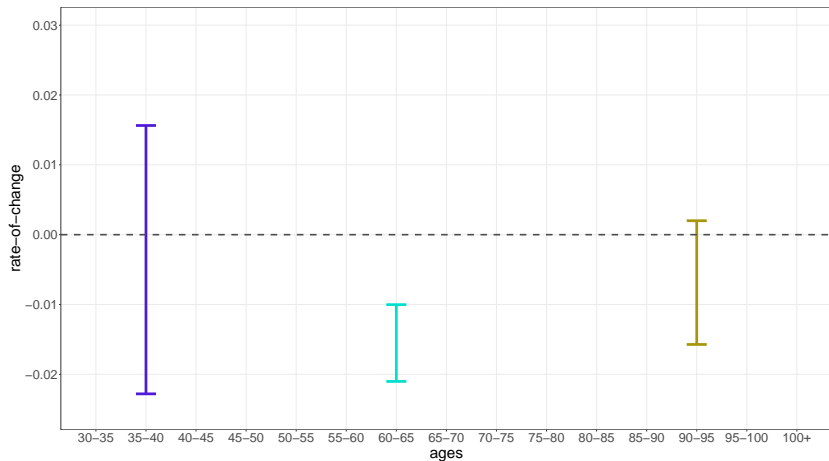
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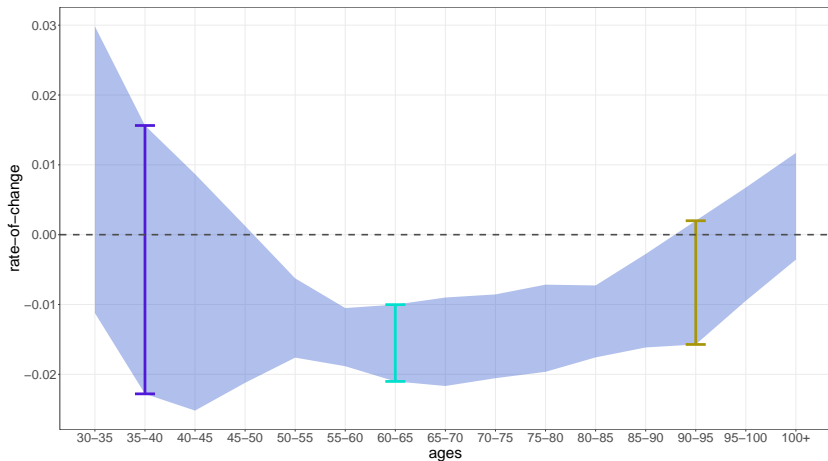
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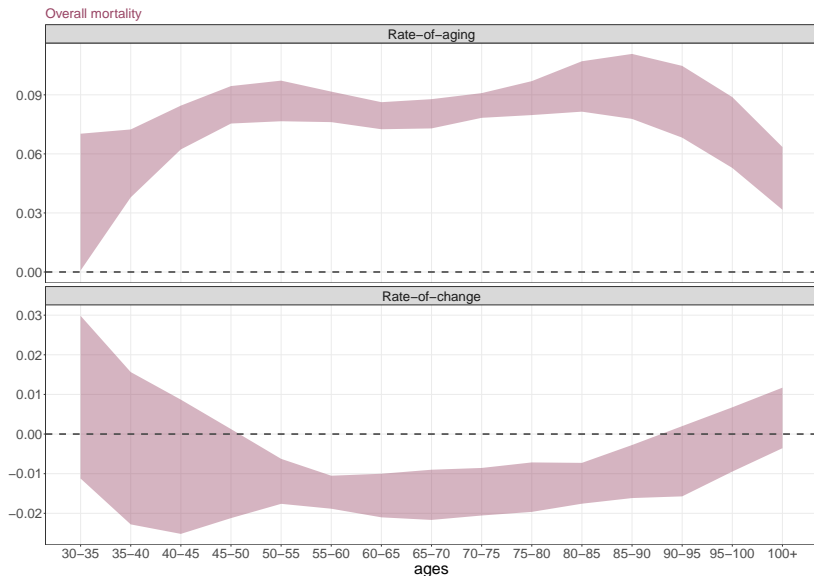
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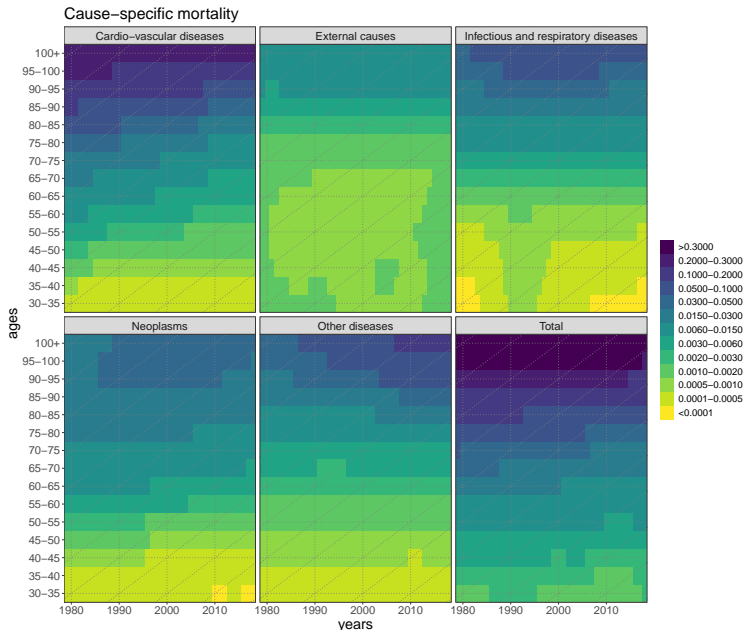


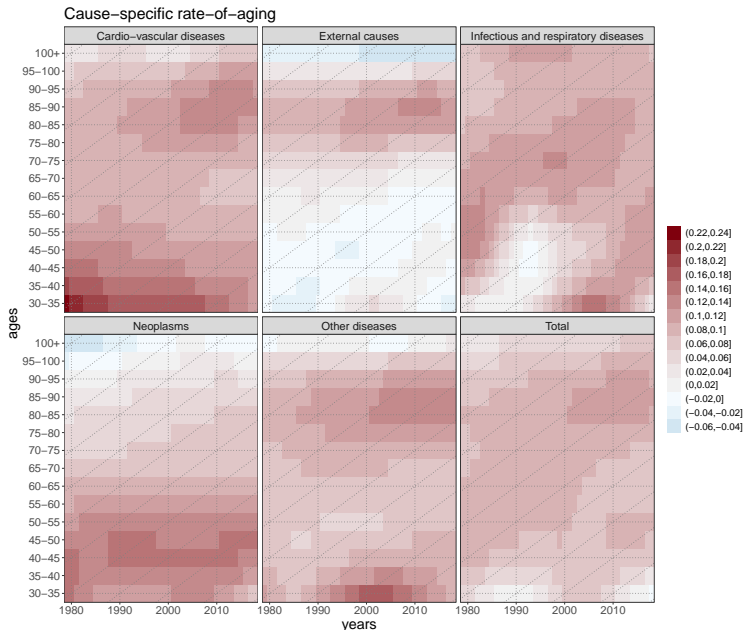




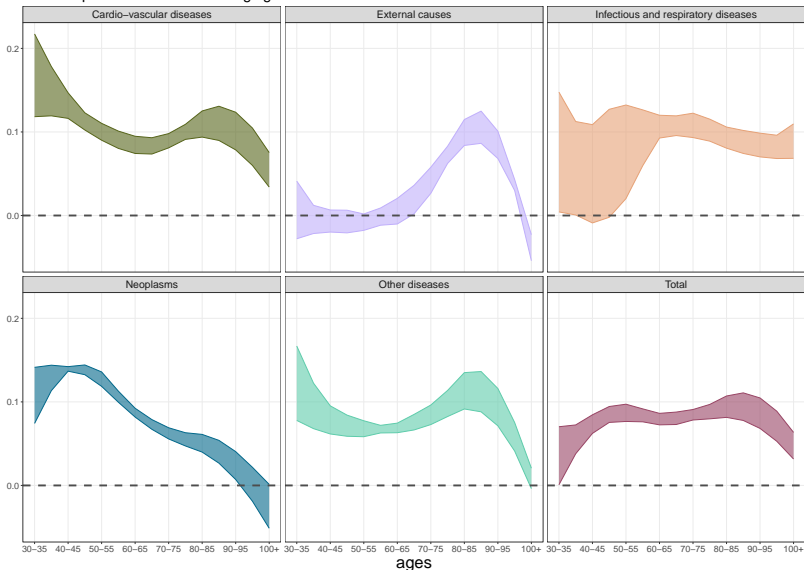


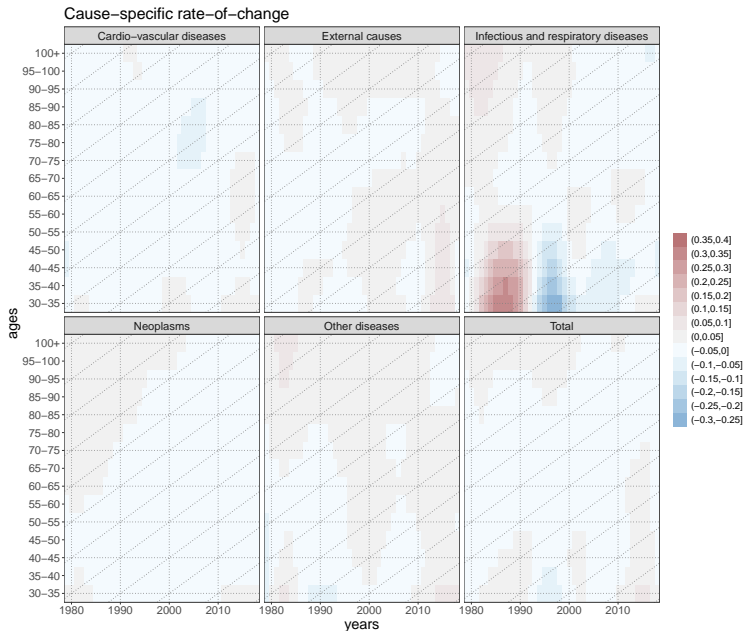
Looking at cause-specific mortality



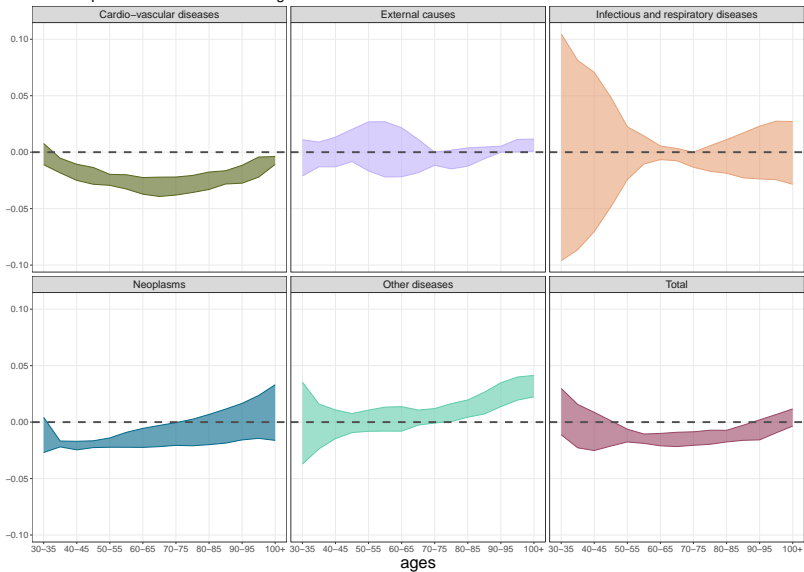


Cause-specific 95% of rate-of-aging





Cause-specific 50% of rate-of-change



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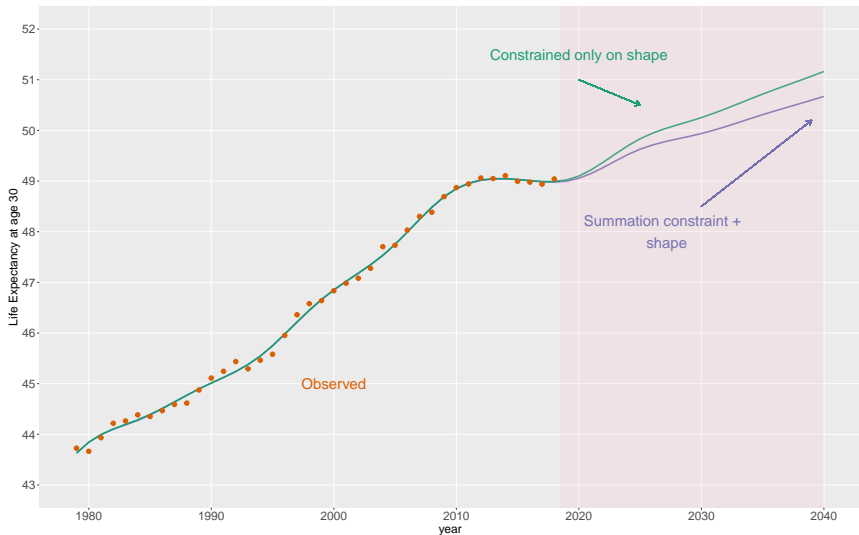
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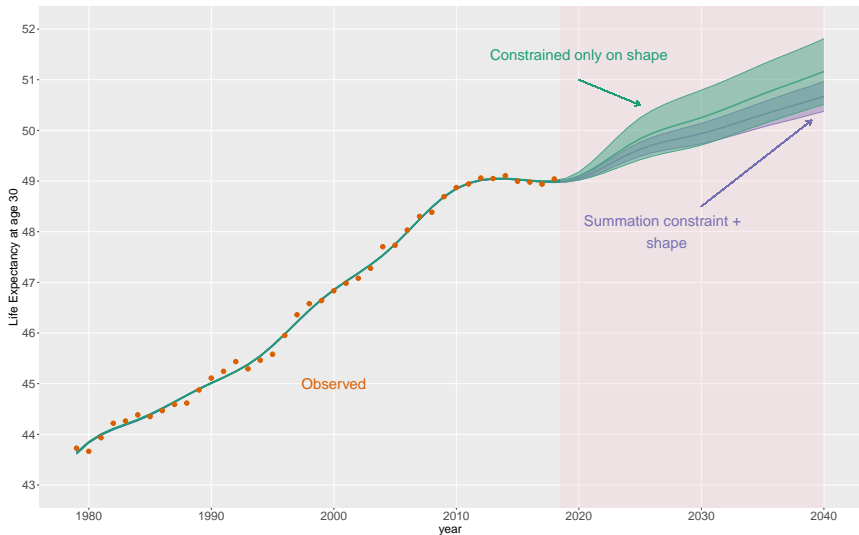
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- We generalize *CP*-splines in a 3D setting (age+time+cause)

Grey lines and areas depict unconstrained model

Life expectancy at age 30 for overall mortality



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- Future research:
 - Selection of the $2 \times (k + 1)$ smoothing parameters
 - Back-testing for model performance
 - Extension to female/male and regional mortality settings

GIANCARLO CAMARDA
in collaboration with MARÍA DURBÁN

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Thanks for your attention.
Comments and questions?



- Camarda, C.G. (2019). Smooth Constrained Mortality Forecasting. *Demographic Research*
- Camarda, C.G. and Durban, M. (2023). Coherent cause-specific mortality forecasting via constrained penalized regression models. *Proc. International Workshop Stat Mod*
- Currie, I.D., Durban, M. and Eilers, P.H.C. (2004). Smoothing and Forecasting Mortality Rates. *Statistical Modelling*
- Currie, I.D., Durban, M. and Eilers, P.H.C. (2006). Generalized Linear Array Models with Applications to Multidimensional Smoothing. *Journal of Royal Statistical Society. Series B*
- Foreman, K.J., Li, G., Best, N. and Ezzati, M. (2017). Small Area Forecasts of Cause-Specific Mortality: Application of a Bayesian Hierarchical Model to US Vital Registration Data. *Journal of the Royal Statistical Society: Series C*
- Kjærgaard, S., Ergemen, Y.E., Kallestrup-Lamb, M. et al. (2019). Forecasting causes of death by using compositional data analysis: the case of cancer deaths. *Journal of the Royal Statistical Society. Series C*
- Wilmoth, J.R. (1995). Are mortality projections always more pessimistic when disaggregated by cause of death? *Mathematical Population Studies*