

# Understanding and assessing climate-driven mortality risk

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# Outline of the presentation

- 1 Background information
- 2 Joint extremes in temperature and mortality
- 3 Excess mortality under climate scenarios
- 4 Cold-related cause-of-death mortality
- 5 The ultimate research questions

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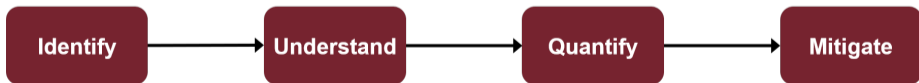
# The link between climate change and human mortality

- Approximately **9** out of every **100** deaths in the world during **2000–2019** were due to extreme cold temperatures. - A recent study
- During **2000–2019**, the mortality burden attributable to extreme temperatures in Australia is estimated to be **11.40%** of the total deaths. - A recent study
- **Between 2030 and 2050**, climate change is expected to cause approximately **250,000** additional deaths per year. - WHO



# What can insurance do in a changing climate?

**Unanticipated** adverse claim experience due to climate change can lead to **insolvency** of insurance and reinsurance companies.



# How can climate change kill you?

According to the WHO:

Between **2030** and **2050**, climate change is expected to cause approximately **250,000** additional deaths **per year**.

Weather-related catastrophes (etc. floods, bushfires and earthquakes)

**Extreme temperatures (etc. heat waves and cold spells)**

Climate-sensitive infectious diseases (etc. malaria disease and vector-borne diseases)

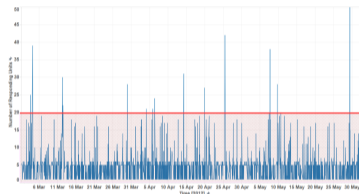
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# The key challenge

A key challenge in modeling extreme risks: scarcity of extreme events. Extreme value theory (EVT) tackles this problem by providing results beyond observed values.

- **Block Maxima:**  
Distribution of the sample maximum
- **Peaks Over Threshold:**  
Distribution of values over a high threshold



**A simplified example:**

**Univariate POT:**

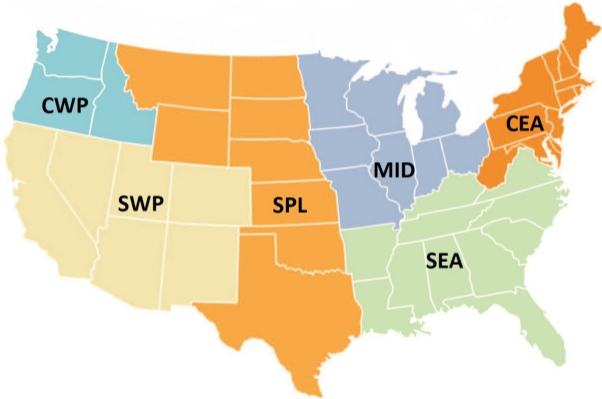
The temperature is  $> 43^\circ$ , how likely is it to be  $> 45^\circ$ ?

**Bivariate POT:**

The temperature is  $> 43^\circ$ , how likely is it that  $> 20$  people die?

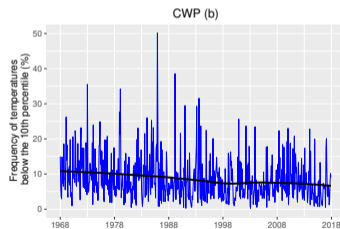
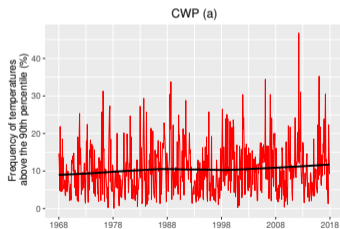
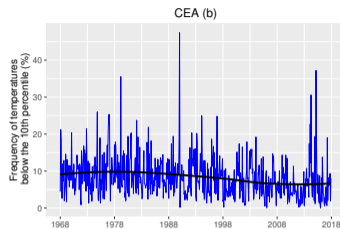
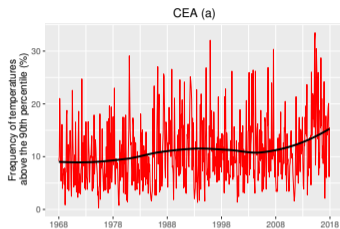


# Data Description and Modeling

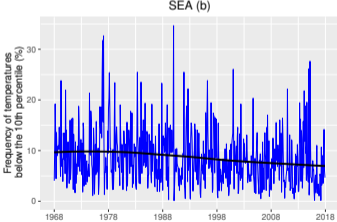
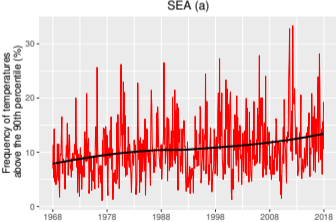
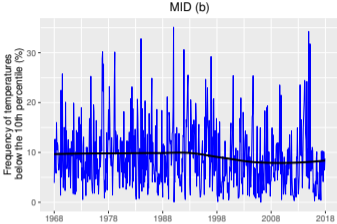
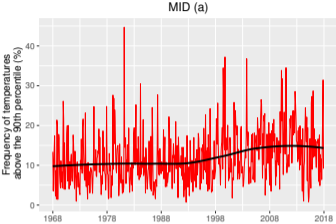


Source: Actuaries Climate Index Executive Summary (2018), Page 4, Figure 2

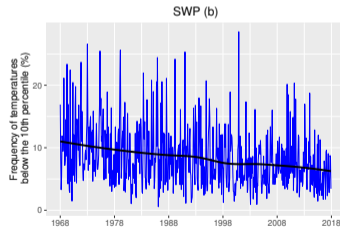
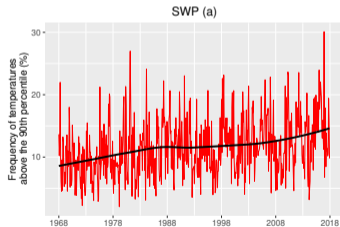
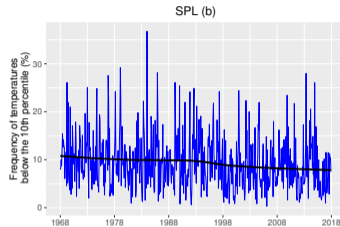
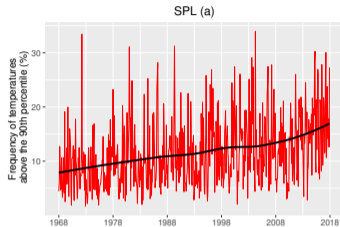
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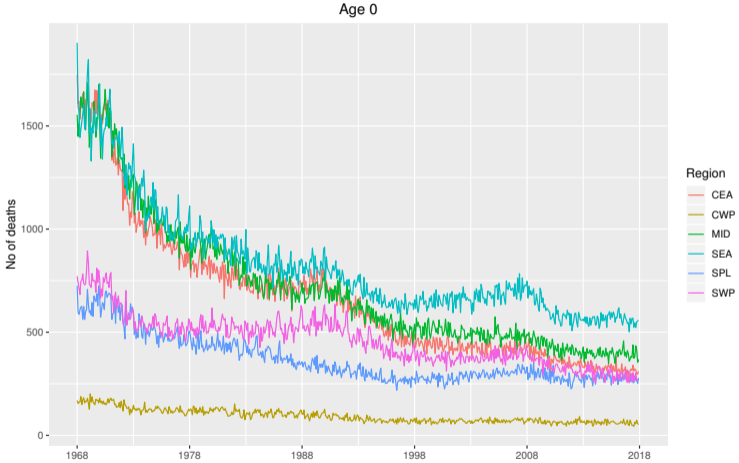
We adopt the **seasonal ARIMA** model which incorporates both non-seasonal and seasonal factors in a multiplicative model, which can be expressed as

$$\text{ARIMA}(p, d, q) \times (P, D, Q)_S, \quad (1)$$

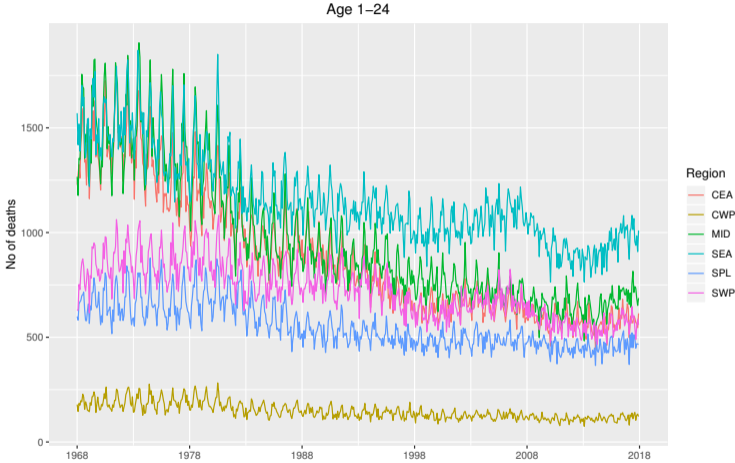
where:

- $p$ ,  $d$ , and  $q$  denote the order of the AR model, the order of differencing, and the order of the MA model in the non-seasonal part, respectively,
- $P$ ,  $D$ , and  $Q$  denote the order of the AR model, the order of differencing, and the order of the MA model in the seasonal part, respectively, and
- $S$  is the time span of repeating the seasonal pattern. Since we are modeling monthly  $T90$  and  $T10$  time series,  $S$  is set to be 12.

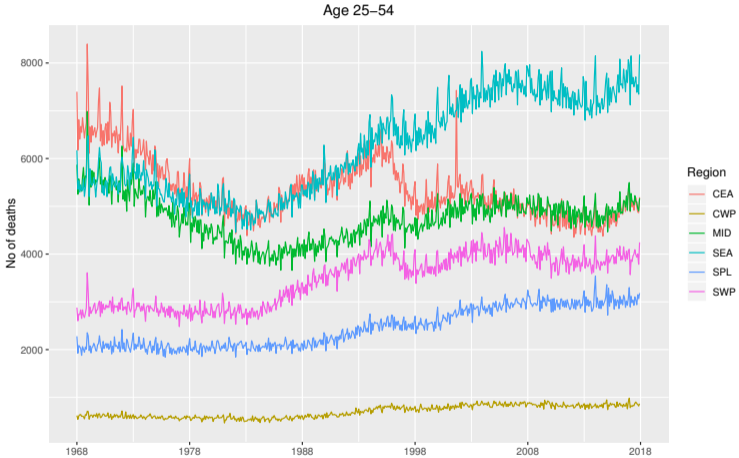
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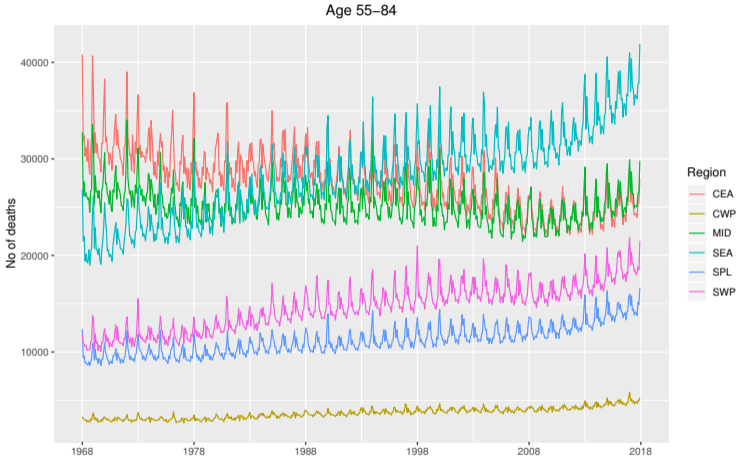


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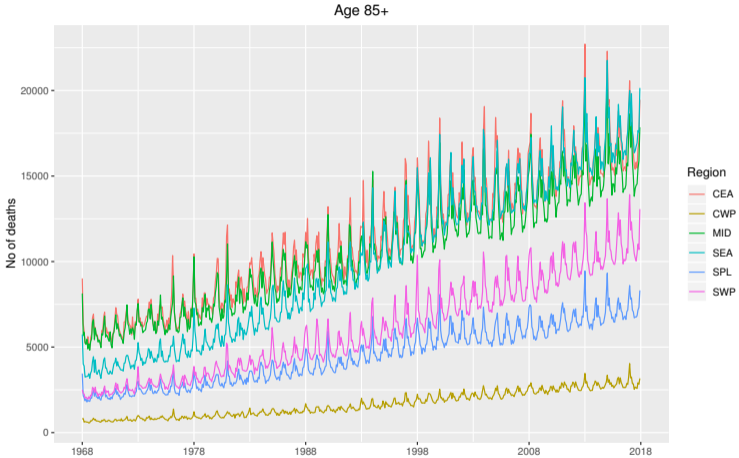




# Data Description and Modeling



# Data Description and Modeling



# Data Description and Modeling

Similar to *T10* and *T90*, we want to obtain the “noise” in death counts via time series models.

- We fit a **seasonal ARIMA** model first.
- We include the **GARCH** component if the resulting residuals fail the Ljung-Box test at 5% level of significance.
- The optimal model is selected based on **AIC**.

# Multivariate Extreme Value Theory

Consider a random variable  $X$  with distribution  $F$  on  $\mathbb{R}$  and denote by  $M_n$  the maximum of a sample of size  $n$  from  $F$ . If there exist norming constants  $a_n > 0$  and  $b_n \in \mathbb{R}$  such that

$$\lim_{n \rightarrow \infty} \Pr \left( \frac{M_n - b_n}{a_n} \leq y \right) = G(y), \quad y \in \mathbb{R}, \quad (2)$$

then we say that  $F$  belongs to the max-domain of attraction of  $G$ , and call  $G$  a **generalized extreme value (GEV)** distribution. The GEV distribution function  $G$  must be of the same type as

$$G(y) = \exp \left\{ - \left( 1 + \gamma \frac{y - \mu}{\sigma} \right)_+^{-1/\gamma} \right\}, \quad (3)$$

where  $\mu \in \mathbb{R}$ ,  $\sigma > 0$ , and  $\gamma \in \mathbb{R}$  are the location, scale, and shape parameters, respectively, and  $c_+ = \max(c, 0)$ .

# Multivariate Extreme Value Theory

Following the works of Balkema and de Haan (1974) and Pickands (1975), the conditional distribution of the normalized exceedance over a high threshold converges to a **generalized Pareto distribution (GPD)**, that is,

$$\lim_{n \rightarrow \infty} \Pr \left( \frac{X - b_n}{a_n} \leq y \mid X > b_n \right) = H(y), \quad y > 0, \quad (4)$$

where  $H$  is of the same type as

$$H(y) = 1 - \left( 1 + \gamma \frac{y - \mu}{\sigma} \right)_+^{-1/\gamma}, \quad (5)$$

with the location, scale, and shape parameters  $\mu \in \mathbb{R}$ ,  $\sigma > 0$ , and  $\gamma \in \mathbb{R}$ . The GPD  $H$  above is supported on the region of  $y$  defined by  $y > 0$  and  $1 + \gamma \frac{y - \mu}{\sigma} > 0$ .

# Multivariate Extreme Value Theory

Consider a  $d$ -dimensional random vector  $\mathbf{X}$  with distribution  $F$  on  $\mathbb{R}^d$  and denote by  $M_n$  the component-wise maximum of a sample of size  $n$  from  $F$ . The limit distribution  $G$ , called a **multivariate GEV distribution**, has marginal distributions  $G_i$  for  $1 \leq i \leq d$  identical to

$$\lim_{n \rightarrow \infty} \Pr \left( \frac{M_n^{(i)} - b_n^{(i)}}{a_n^{(i)}} \leq y \right) = G_i(y), \quad (6)$$

which therefore is of the same type as Equation (5).

In practice, it is common to first transform the marginal distributions to a particular distribution before fitting a multivariate GEV distribution. In this paper, we choose the unit Fréchet transformation

$$z = -\frac{1}{\log G_i(y)}. \quad (7)$$

# Multivariate Extreme Value Theory

According to Propositions 5.10 and 5.11 in Resnick (1987), the representation of a multivariate GEV distribution with unit Fréchet margins can be written as

$$G(\mathbf{y}) = \exp\{-V(\mathbf{z})\}, \quad (8)$$

where  $V(\cdot)$ , the exponent measure, has a functional representation

$$V(\mathbf{z}) = \int_{S_d} \max_{1 \leq i \leq d} \left( \frac{q_i}{z_i} \right) d\phi(\mathbf{q}), \quad (9)$$

with  $\phi$  being a finite spectral measure on  $S_d = \{\mathbf{q} \in \mathbb{R}^d : \|\mathbf{q}\| = 1\}$ , and  $\|\cdot\|$  representing an arbitrary norm in  $\mathbb{R}^d$ . We also impose a constraint such that, for  $1 \leq i \leq d$ ,

$$\int_{S_d} q_i d\phi(q_i) = 1, \quad (10)$$

but beyond this the spectral measure  $\phi$  is unknown.

# Multivariate Extreme Value Theory

As in this study we focus on assessing the upper tail dependence between temperature and mortality, we adopt the **symmetric logistic model** for the function  $V$ , which is a natural candidate and a commonly used dependence model in bivariate POT studies (see *e.g.* Tawn, 1990; Coles *et al.*, 1999; Rootzén and Tajvidi, 2006). Under the symmetric logistic model,

$$V(z_1, z_2) = (z_1^{-r} + z_2^{-r})^{1/r}, \quad r \geq 1, \quad (11)$$

which can be retrieved from Equation (9) with a suitably chosen spectral measure  $\phi$  on  $S_2$ . The exponent measure  $V(z_1, z_2)$  determines the strength of dependence between the two margins. **In particular, independence is obtained when  $r = 1$ , and perfect dependence is obtained as  $r \rightarrow \infty$ .**



# Multivariate Extreme Value Theory

The multivariate POT theorem then states that, for a random vector  $\mathbf{X}$  distributed by  $F \in \text{MDA}(G)$ , assuming  $0 < G(0) < 1$  without loss of generality, the conditional distribution of  $\mathbf{a}_n^{-1}(\mathbf{X} - \mathbf{b}_n)$  given  $\mathbf{X} \not\leq \mathbf{b}_n$  converges to the multivariate GPD as

$$H(\mathbf{y}) = \frac{1}{-\log G(0)} \log \frac{G(\mathbf{y})}{G(\mathbf{y} \wedge \mathbf{0})}, \quad (12)$$

which is defined for all  $\mathbf{y} \in \mathbb{R}^d$  such that  $G(\mathbf{y}) > 0$ . In particular,  $H(\mathbf{y}) = 0$  for  $\mathbf{y} < 0$  and  $H(\mathbf{y}) = 1 - \frac{\log G(\mathbf{y})}{\log G(0)}$  for  $\mathbf{y} > 0$  (Rootzén and Tajvidi, 2006; Rootzén *et al.*, 2018a,b).

# Empirical Results

The Pickands dependence function  $A : [0, 1] \rightarrow [0, 1]$  is defined as

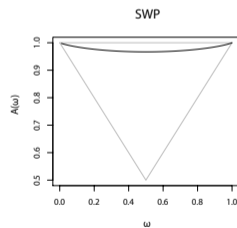
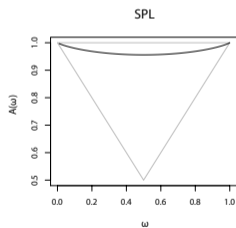
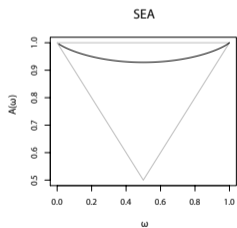
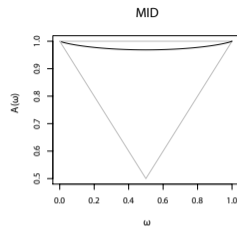
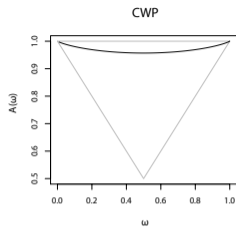
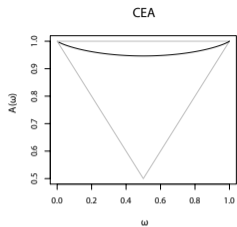
$$A(\omega) = \int_{S_2} \max(\omega q_1, (1 - \omega)q_2) d\phi(\mathbf{q}), \quad 0 \leq \omega \leq 1, \quad (13)$$

which links the function  $V$  through the relation

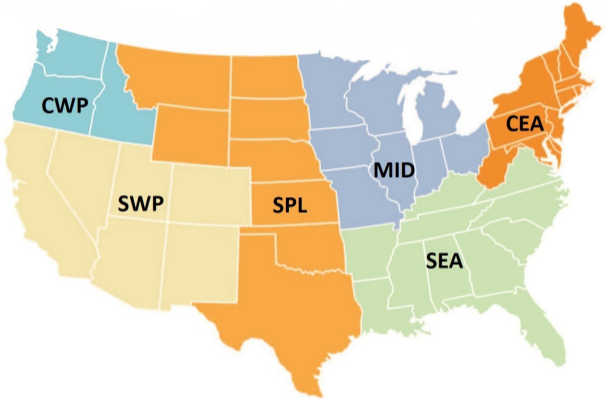
$$A(\omega) = \frac{V(z_1, z_2)}{z_1^{-1} + z_2^{-1}}, \quad (14)$$

with  $\omega = \frac{z_2}{z_1 + z_2}$ . By Equation (10), it is clear that  $A(0) = A(1) = 1$ . **If two random variables with unit Fréchet margins are independent, then  $A(\omega) = 1$  for all  $0 \leq \omega \leq 1$ , while if they are perfectly dependent, then  $A(\omega) = \max(\omega, 1 - \omega)$  for all  $0 \leq \omega \leq 1$ .**

# Empirical Results



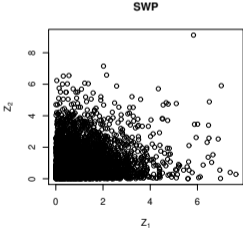
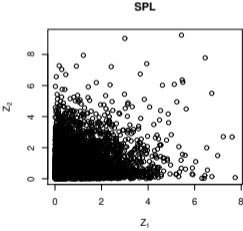
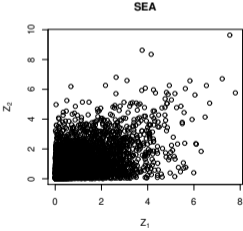
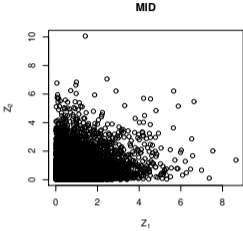
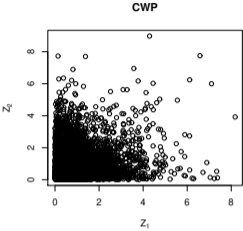
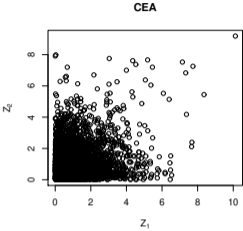
# Empirical Results



Source: Actuaries Climate Index Executive Summary (2018), Page 4, Figure 2



# Empirical Results



# Conclusions

Li, H., Tang, Q., 2022. Joint extremes in temperature and mortality: A bivariate POT approach. North American Actuarial Journal, 26(1), 43–63.

Frequency of extreme hot temperatures  
→ weak impact on death counts

Heatwaves?? A heatwave is generally defined in terms of a consecutive period of excessively hot weather (4 days)

“Harvesting effect” or “Mortality displacement”

A simple measure of monthly hot temperature frequency may not be adequate

Frequency of extreme cold weather  
→ stronger impact on older people aged 55–84 and 85+

Cold weather can cause substantial short-term increase in mortality

Epidemics of influenza are likely to be associated with extreme cold weather

The increase in mortality following extreme cold is long lasting

**The elderly are more fragile to extreme temperatures**

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# Data Description and Modeling





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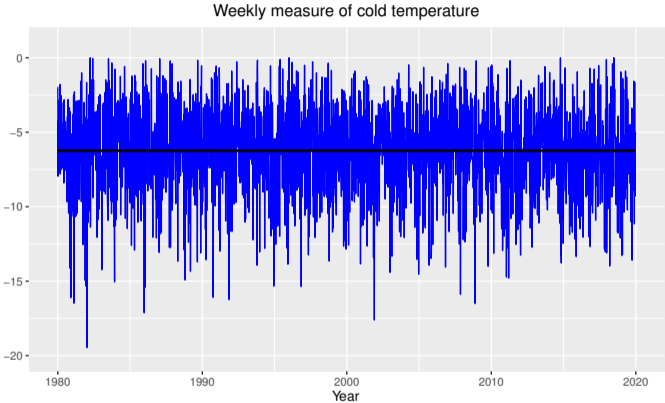


Figure 1: Weekly measures of extreme cold temperature

# Data Description and Modeling

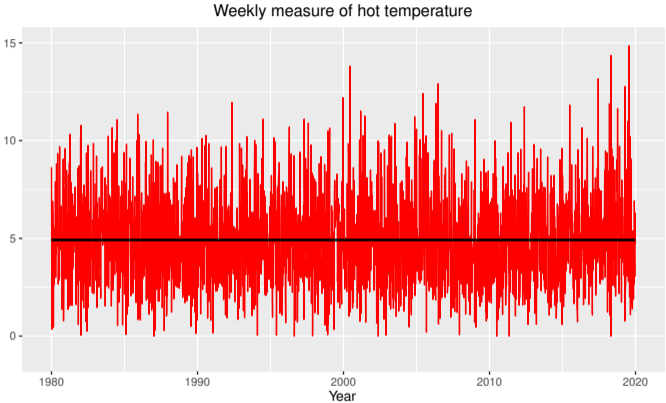


Figure 2: Weekly measures of extreme hot temperature

# Data Description and Modeling

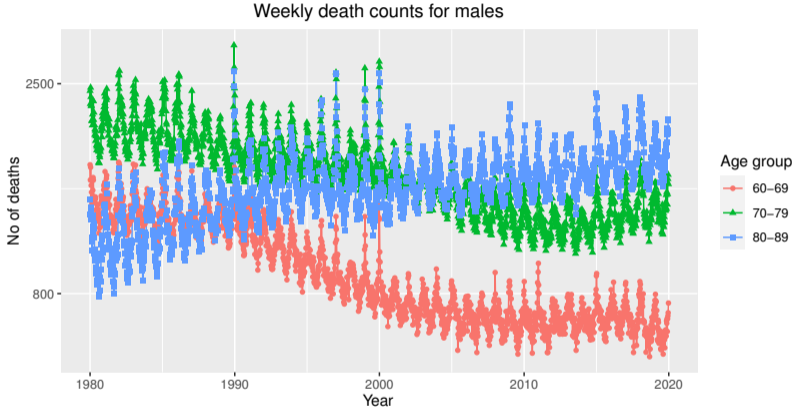


Figure 3: Weekly death counts for England & Wales: Males



# Data Description and Modeling

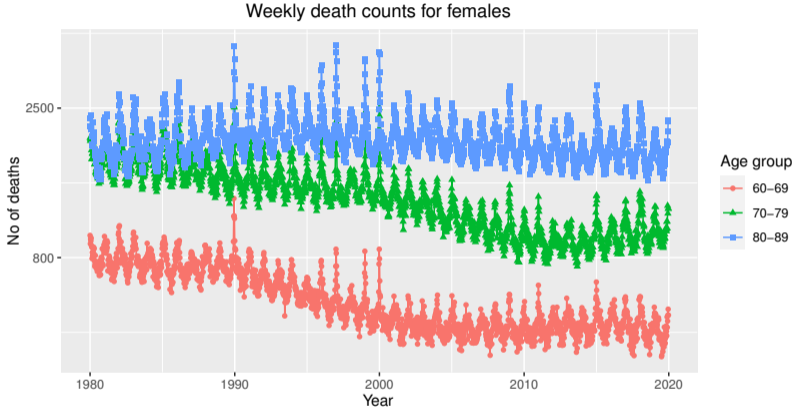


Figure 4: Weekly death counts for England & Wales: Females

# Empirical Results

Coles *et al.* (1999) developed the index  $\chi$  to measure extreme dependence for bivariate random variables. Assuming that random variables  $Z_1$  and  $Z_2$  have the same marginal distribution  $F$ , the index  $\chi$  is defined as

$$\chi = \lim_{u \uparrow 1} \Pr(F(Z_2) > u | F(Z_1) > u). \quad (15)$$

Thus,  $\chi$  denotes the probability of one variable reaching the extreme value given that the other variable has already reached it. **If  $\chi = 0$ , the two variables are said to be asymptotically independent. While for full tail dependence, we have  $\chi = 1$ .**

# Empirical Results

**Table 1:** Extreme dependence measure  $\chi$  based on the bivariate POT analysis

Male												
Lag	0			1			2			3		
Age	60–69	70–79	80–89	60–69	70–79	80–89	60–69	70–79	80–89	60–69	70–79	80–89
<b>Tmin</b>	0.000	0.001	0.001	<b>0.087</b>	<b>0.091</b>	<b>0.132</b>	0.037	0.039	0.018	0.014	0.012	0.001
<b>Tmax</b>	<b>0.059</b>	<b>0.076</b>	<b>0.148</b>	0.000	0.000	0.000	0.000	0.000	0.001	0.001	0.000	0.001
Female												
Lag	0			1			2			3		
Age	60–69	70–79	80–89	60–69	70–79	80–89	60–69	70–79	80–89	60–69	70–79	80–89
<b>Tmin</b>	0.000	0.001	0.001	<b>0.055</b>	<b>0.092</b>	<b>0.129</b>	0.019	0.039	0.027	0.000	0.001	0.011
<b>Tmax</b>	<b>0.072</b>	<b>0.080</b>	<b>0.130</b>	0.001	0.001	0.001	0.001	0.001	0.001	0.001	0.001	0.001

# Empirical Results

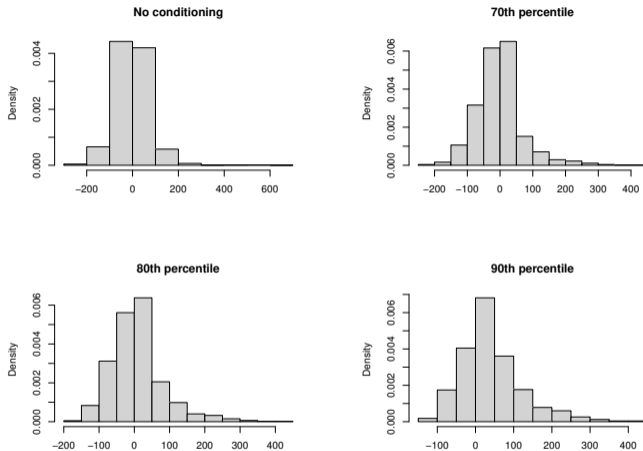


Figure 5: Density functions of excess deaths under different  $T_{cold}$  scenarios: Males, 80–89

# Empirical Results

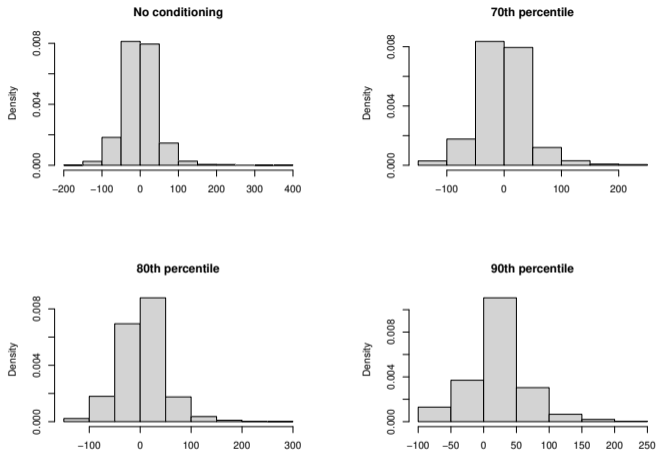


Figure 6: Density functions of excess deaths under different  $T_{cold}$  scenarios: Males, 60–69



## What did we find?

A, Chaudhry., M, Leitschkis., Li, H., Tang, Q., 2023. An EVT Approach to Quantifying Mortality Risk of Extreme Temperatures.

Frequency of extreme hot temperatures  
→ impact on older people

Short-term increase in mortality happens immediately after extreme hot temperature events.

Consistent with the impact of heatwaves.

Frequency of extreme cold weather  
→ impact on older people

The increase in mortality following extreme cold is longer lasting

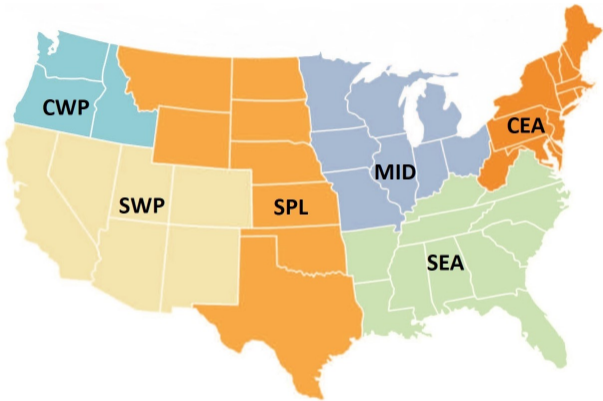
Short-term increase in mortality happens one week after extreme cold temperature events.

**The “oldest old” are more fragile to extreme temperatures**

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# T10 index for extreme cold temperatures

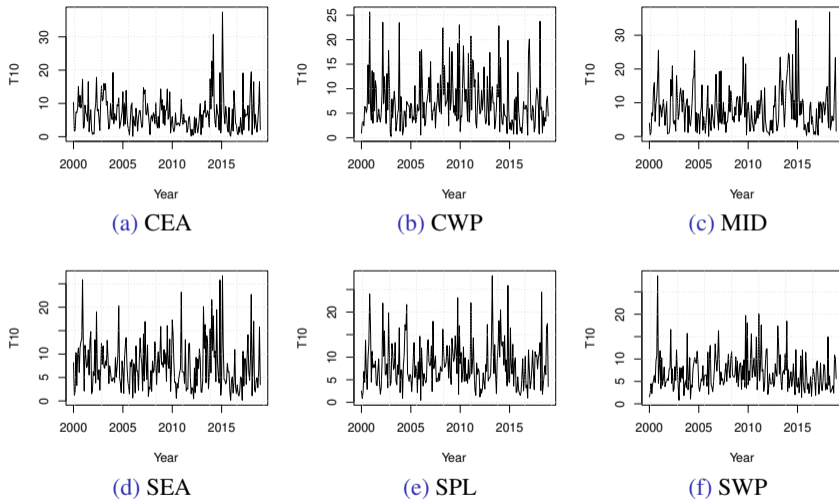
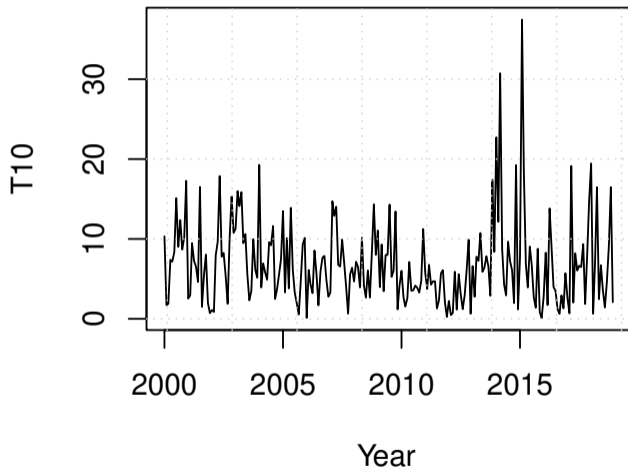


Figure 7: T10 index over 2000–2018.

## T10 index - Central East Atlantic (CEA)



# Cause-of-death definitions based on International Codification of Diseases

Table 2: Codification of five major causes of death

Cause of death	ICD-10 code
Diabetes	E10–E14
External	V01–Y89
Respiratory	J09–J98
Neoplasms	C00–D48
Vascular	I00–I78

# Cause of death data

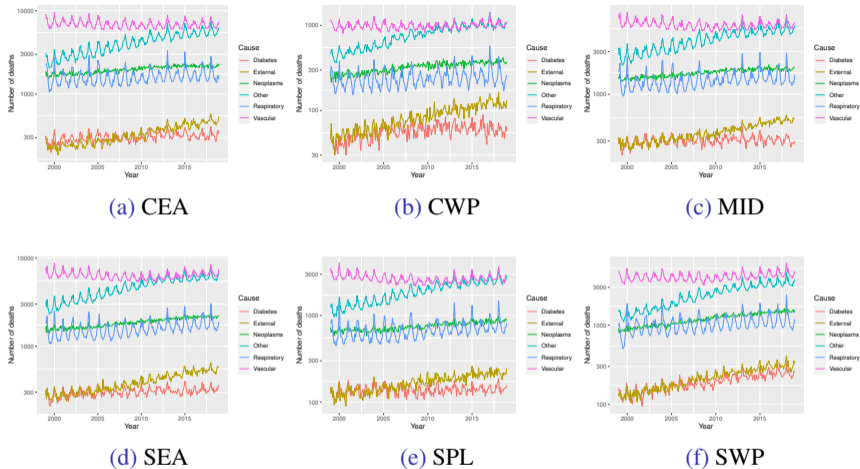


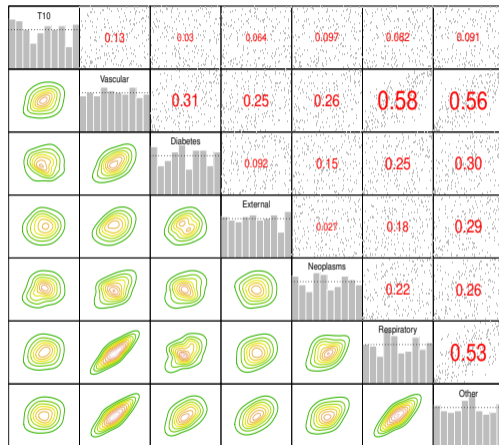
Figure 8: Monthly death counts for ages 85+.

# Cause of death data - Central East Atlantic (CEA)





# Pairwise dependence structure



# A quick introduction to Vine-copula modeling

Key idea: is to construct a **flexible** dependence structure across variables using **pair-copulas** as **bivariate building blocks** Aas *et al.* (2009).

Consider a simplified case with three causes of death, namely  $A$ ,  $B$ , and  $C$ . Under the vine copula framework, the joint probability distribution  $f_{ABC}$  can be expressed as follows

$$f_{ABC} = C_{AC} \times C_{BC} \times C_{AB|C} \times f_A \times f_B \times f_C, \quad (16)$$

where  $C$  denotes bivariate pair-copulas and  $f$  denotes marginal distributions. As such, the joint density of excess deaths is broken down into a product of **bivariate copulas** and **marginal densities**.

## A quick introduction to vine copula modeling

$AC$  could be assigned a copula with **upper tail dependence** (e.g. Gumbel),  $BC$  could be assigned a copula with **lower tail dependence** (e.g. Clayton) and  $AB|C$  could be assigned a copula with **no tail dependence** (e.g. Gaussian).

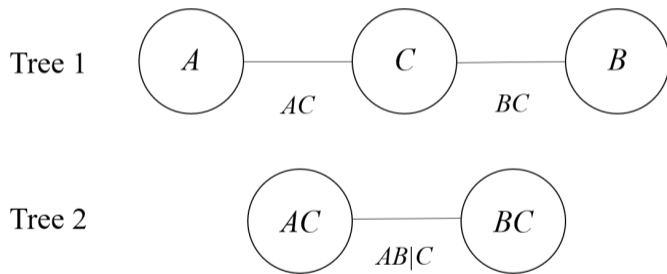


Figure 9: Example of an R-vine tree sequence

# Three different climate scenarios

In particular, we analyze the distribution of excess deaths at time  $t$  under three temperature scenarios as follows

- 1 The  $T10$  index at time  $t$  exceeds its 90th percentile.
- 2 The  $T10$  index at time  $t - 1$  exceeds its 90th percentile.
- 3 The  $T10$  index at both time  $t$  and  $t - 1$  exceeds its 90th percentile.

**Table 3:** Excess deaths breakdown by causes: Scenario 1

Region	Diabetes	External	Neoplasms	Other	Respiratory	Vascular
CEA	1.3%	1.1%	3.9%	28.2%	25.3%	<b>40.2%</b>
CWP	1.3%	3.7%	-1.3%	<b>40.5%</b>	30.2%	25.6%
MID	0.6%	0.9%	-2.6%	14.8%	<b>61.8%</b>	24.4%
SEA	1.2%	2.8%	4.0%	40.1%	11.3%	<b>40.7%</b>
SPL	0.4%	5.7%	4.5%	<b>38.0%</b>	23.0%	28.3%
SWP	1.8%	1.0%	3.3%	<b>44.4%</b>	12.2%	37.3%

Table 4: Excess deaths breakdown by causes: Scenario 2

Region	Diabetes	External	Neoplasms	Other	Respiratory	Vascular
CEA	8.0%	-7.5%	-6.6%	45.5%	<b>47.3%</b>	13.2%
CWP	3.2%	3.3%	-1.0%	35.2%	<b>39.2%</b>	20.1%
MID	-0.6%	-0.3%	-1.9%	-4.2%	<b>117.4%</b>	-10.5%
SEA	0.3%	0.5%	-1.5%	37.4%	23.4%	<b>39.9%</b>
SPL	1.6%	2.7%	-1.8%	-30.9%	40.2%	<b>88.2%</b>
SWP	9.5%	0.4%	-26.4%	<b>102.0%</b>	43.3%	-28.8%

**Table 5:** Excess deaths breakdown by causes: Scenario 3

Region	Diabetes	External	Neoplasms	Other	Respiratory	Vascular
CEA	2.1%	-0.1%	2.0%	29.9%	32.0%	<b>34.0%</b>
CWP	1.9%	3.4%	-0.9%	35.5%	<b>36.2%</b>	24.0%
MID	0.1%	0.4%	-2.8%	8.3%	<b>81.0%</b>	12.9%
SEA	0.9%	1.9%	2.2%	37.7%	16.8%	<b>40.5%</b>
SPL	1.0%	4.5%	2.3%	15.7%	27.3%	<b>49.1%</b>
SWP	1.2%	0.2%	0.7%	<b>65.5%</b>	13.5%	19.0%

# Empirical results

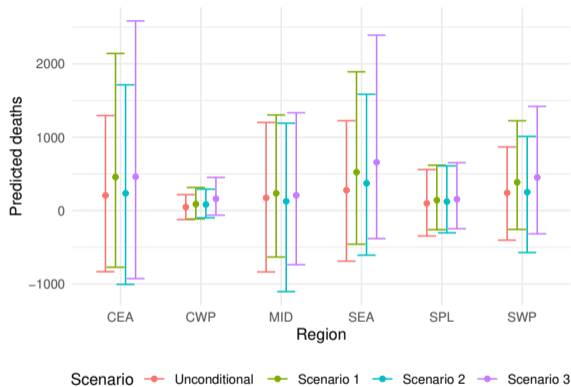
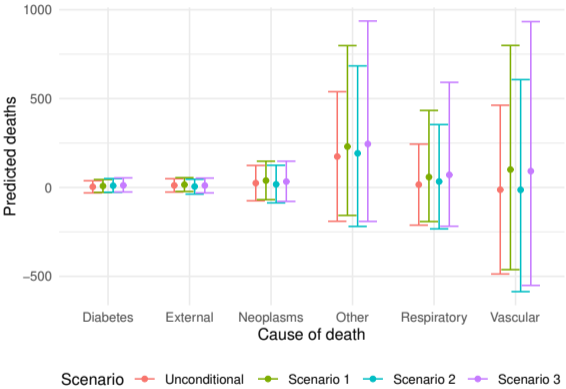


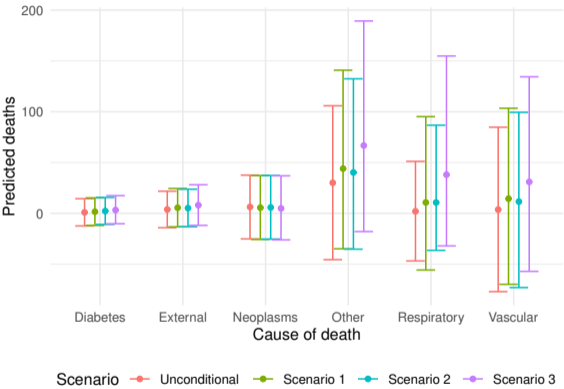
Figure 10: Prediction intervals of monthly total deaths at 10th, 50th, and 90th percentiles.



# Empirical results



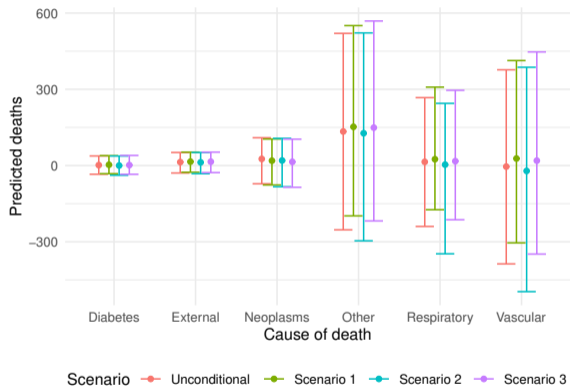
(a) CEA



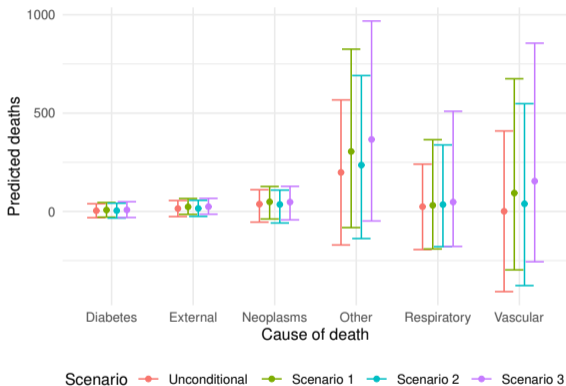
(b) CWP



# Empirical Results

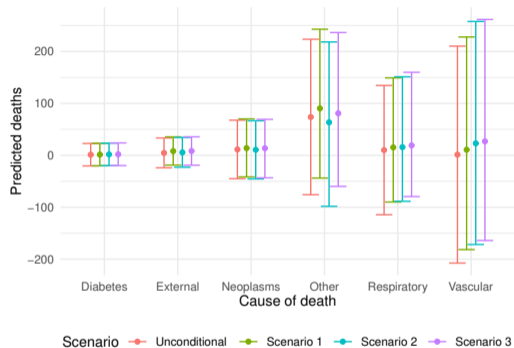


(c) MID

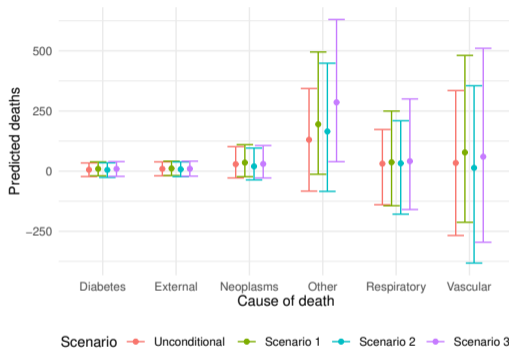


(d) SEA

# Empirical Results



(e) SPL



(f) SWP

Figure 11: Prediction intervals of monthly cause-specific deaths at 10th, 50th, and 90th percentiles.

# Empirical results

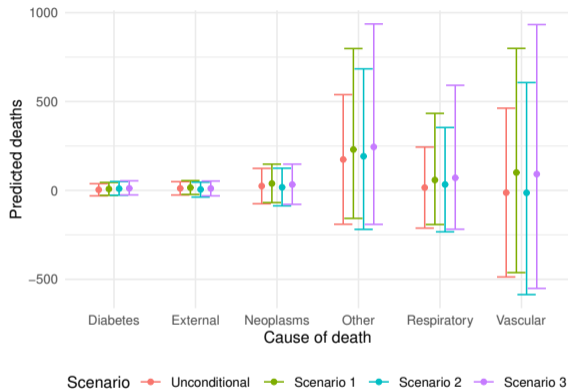


Figure 12: Prediction intervals of monthly total deaths at 10th, 50th, and 90th percentiles for CEA.

# Outline

- 1 Background information
- 2 Joint extremes in temperature and mortality
- 3 Excess mortality under climate scenarios
- 4 Cold-related cause-of-death mortality
- 5 The ultimate research questions**

# The ultimate research questions

Our result is expected to provide insights into the following questions:

- ① Who are the excess deaths? – Find the age groups that are particularly sensitive to climate change.
- ② When do excess deaths occur? – Determine if more excess deaths occur in winter or summer.
- ③ Where are the excess deaths? – Identify regions that are most vulnerable to climate change.

## Questions and discussions



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